Joint association of urinary sodium and potassium excretion with cardiovascular events and mortality: prospective cohort study

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ABSTRACT

OBJECTIVE
To evaluate the joint association of sodium and potassium urinary excretion (as surrogate measures of intake) with cardiovascular events and mortality, in the context of current World Health Organization recommendations for daily intake (<2.0 g sodium, >3.5 g potassium) in adults.

DESIGN
International prospective cohort study.

SETTING
18 high, middle, and low income countries, sampled from urban and rural communities.

PARTICIPANTS
103 570 people who provided morning fasting urine samples.

MAIN OUTCOME MEASURES
Association of estimated 24 hour urinary sodium and potassium excretion (surrogates for intake) with all cause mortality and major cardiovascular events, using multivariable Cox regression. A six category variable for joint sodium and potassium was generated: sodium excretion (low (<3 g/day), moderate (3.5 g/day), and high (>5 g/day) sodium intakes) by potassium excretion (greater/equal or less than median 2.1 g/day).

RESULTS
Mean estimated sodium and potassium urinary excretion were 4.93 g/day and 2.12 g/day, respectively. After a median follow-up of 8.2 years, 7884 (6.1%) participants had died or experienced a major cardiovascular event. Increasing urinary sodium excretion was positively associated with increasing potassium excretion (unadjusted r=0.34), and only 0.002% had a concomitant urinary excretion of <2.0 g/day of sodium and >3.5 g/day of potassium. A J-shaped association was observed of sodium excretion and inverse association of potassium excretion with death and cardiovascular events. For joint sodium and potassium excretion categories, the lowest risk of death and cardiovascular events occurred in the group with moderate sodium excretion (3-5 g/day) and higher potassium excretion (21.9% of cohort). Compared with this reference group, the combinations of low potassium with low sodium excretion (hazard ratio 1.23, 1.11 to 1.37; 7.4% of cohort) and low potassium with high sodium excretion (1.10, 1.02 to 1.18; 29.6% of cohort) were associated with the highest risk, followed by low sodium excretion (1.19, 1.02 to 1.38; 3.3% of cohort) and high sodium excretion (1.10, 1.02 to 1.18; 29.6% of cohort) among those with potassium excretion greater than the median. Higher potassium excretion attenuated the increased cardiovascular risk associated with high sodium excretion (P for interaction=0.007).

CONCLUSIONS
These findings suggest that the simultaneous target of low sodium intake (<2 g/day) with high potassium intake (>3.5 g/day) is extremely uncommon. Combined moderate sodium intake (3-5 g/day) with high potassium intake is associated with the lowest risk of mortality and cardiovascular events.

Introduction
Consumption of sodium and potassium is essential to health, as neither are produced endogenously and both are necessary for critical physiological processes.1-3 Maintaining homeostasis requires a...
mutual interdependence of sodium and potassium, meaning that their relation with health is necessarily linked.\textsuperscript{1,3} Despite these fundamental physiological considerations, current public health policy adopts markedly opposing guidance for sodium and potassium intake, with the World Health Organization recommending an extremely low sodium intake (<2 g/day, whereas the current average intake is about 4 g/day) but high potassium intake (>3.5 g/day, whereas the average intake is about 2 g/day) in the entire adult population.\textsuperscript{4,6} International data on the percentage of people with very low intakes of sodium and simultaneously high potassium intakes\textsuperscript{7} and the joint effects on health outcomes from large international epidemiological studies are limited.\textsuperscript{4-5,8-9}

Current public health policy on targets for sodium intake is based primarily on small, and mostly short term, clinical trials evaluating the relation of changes in sodium and potassium intake with blood pressure.\textsuperscript{5-8} Although clinical trials have reported a decrease in blood pressure with a reduction in sodium intake,\textsuperscript{8} many prospective cohort studies have reported a J-shaped association of sodium intake and cardiovascular disease or mortality,\textsuperscript{10-16} with an increased risk emerging at sodium intakes <2.7 g/day and >5 g/day.\textsuperscript{10} The increased cardiovascular risk associated with high sodium intake seems to be especially in those with hypertension,\textsuperscript{11} whereas the increased risk associated with low sodium intake might be mediated through activation of physiological systems to conserve sodium (eg, renin-angiotensin-aldosterone system).\textsuperscript{18} In contrast, the association of potassium intake with blood pressure and cardiovascular disease is consistent, with a linear reduction in blood pressure and cardiovascular risk with increasing potassium intake reported in most studies.\textsuperscript{3} Some epidemiological studies report an interactive association of sodium and potassium intake with blood pressure, and other studies report use of sodium-potassium ratios for predicting risk of cardiovascular events and mortality.\textsuperscript{19-25} Whether a potential modifying effect of increased potassium intake is directly causal or simply a marker of improved diet quality,\textsuperscript{26} as diets rich in fruit and vegetables (which are associated with lower cardiovascular risk) are high in potassium, is uncertain.

In the Prospective Urban Rural Epidemiological (PURE) Study, we have previously reported on the intermediate term (3.7 years of follow-up)\textsuperscript{16,17} association of sodium and potassium intake with mortality and cardiovascular events. In the current analyses, we further explore the association of sodium and potassium intake with death and cardiovascular disease in extended follow-up (eight years).

Methods
Study population
The PURE Study is a large scale epidemiological cohort study that has enrolled people (aged 35 to 70 years) residing in 628 urban and rural communities in low, middle, and high income countries (Bangladesh, India, Pakistan, Zimbabwe, Argentina, Brazil, Chile, Malaysia, Poland, South Africa, Turkey, China, Colombia, Iran, Canada, Sweden, Palestinian territories occupied by Israel, and United Arab Emirates).\textsuperscript{17-29} The supplementary appendix describes the selection of participants. Recruitment began in January 2003. Participants provided written informed consent. For the current analysis, we included 103,570 participants from 18 countries who collected early morning fasting urine samples suitable for analysis, of which 103,200 (99.6%) had follow-up information available (95% completed at least one follow-up visit). The Population Health Research Institute of Hamilton Health Sciences, Ontario, Canada coordinated the study.

Exposure assessments
A morning fasting midstream urine sample, collected from each participant, was frozen at −20°C to −70°C and shipped to a central laboratory in Canada, China, India, or Turkey. The samples were shipped at ambient temperature using the STP 250 ambient specimen shipping box (see supplementary appendix for a description of the methods used for performing urinary analyses). We used the Kawasaki formula to estimate 24 hour urinary sodium and potassium excretion from a fasting morning sample, and these estimates have been validated previously versus measured with 24 hour urine collection.\textsuperscript{30-33} An international validation study reported an intraclass correlation of Kawasaki formula to actual 24 hour urine collections of 0.71.\textsuperscript{31} Therefore, we used the estimates derived from fasting morning urine as surrogates for sodium and potassium intake in the study. (The supplementary appendix provides summary details of validation of this approach.)

Ascertainment of outcomes
Standardised case report forms were used to capture major cardiovascular events (myocardial infarction, stroke, and heart failure), cancer, and death at follow-up, which were adjudicated using standardised definitions.\textsuperscript{32} For the current analysis, we included all adjudicated outcome events in the PURE database up to September 2017. The primary composite outcome was all cause mortality or myocardial infarction or stroke or heart failure. Secondary outcomes were individual components of the primary composite outcome, and new diagnosis of cancer on follow-up.

Ascertainment of covariates
All participants completed baseline standardised questionnaires, which included detailed information on age, sex, lifestyle risk factors (eg, smoking, diet, physical activity, alcohol intake), comorbidities (eg, history of diabetes, hypertension, cardiovascular disease), physical measurements (blood pressure, heart rate, body mass index (BMI), waist:hip ratio), and laboratory measurement of lipoproteins (88%). A modified alternative healthy eating index (mAH EI) score was used to measure overall diet quality, with higher scores indicating a healthier diet.\textsuperscript{33}
Statistical analyses
Baseline differences in characteristics between study participants in different categories of estimated sodium and potassium excretion have been reported previously (see supplementary appendix).\textsuperscript{25} Restricted cubic spline plots were used to explore the shape of association between estimated sodium and potassium excretion and outcomes, fitting a restricted cubic spline function with four knots (5th, 35th, 65th, and 95th centiles).\textsuperscript{26} Our primary outcome measure was the composite of all cause mortality and major cardiovascular events (cardiovascular death, stroke, myocardial infarction, and heart failure). Secondary outcomes included individual components of the primary outcome.

Based on our restricted cubic spline plots for the primary outcome, and the results of previous analyses,\textsuperscript{13,16} we selected 4.00 to 4.99 g/day as the reference category for sodium excretion and <1.5 g/day for potassium excretion. We modelled the time to event outcome using multivariable Cox regression proportional hazards analysis with mixed effect (centre included as a random effect variable to account for clustering by research site) to determine the association between estimated urinary sodium and potassium excretion and outcomes, using three sequential models. Model 1 (primary model) adjusted for age (included as spline function), sex, education, current and former alcohol intake (units weekly), diabetes mellitus, BMI, physical activity, history of cardiovascular events, use of cardiovascular drugs (blood pressure lowering, statins, or antidiabetics), history of tuberculosis, cancer, HIV, and current and former smoking, which constituted our primary multivariable model (with an additional model that included low density lipoprotein (LDL) cholesterol:high density lipoprotein (HDL) cholesterol ratio). Model 2 also included caloric intake, estimated potassium (or sodium) excretion, waist:hip ratio, and modified mAHEI. Model 3 also included history of hypertension, and baseline systolic and diastolic blood pressure and heart rate, which are in the putative causal pathway. We performed an analysis of the combined effects of sodium and potassium, in which we generated six categories for sodium excretion (low (<3 g/day), moderate (3-5 g/day), and high (>5 g/day)), and potassium excretion above and below the median (2.1 g/day). Three categories were selected for sodium excretion, as the estimated association was J-shaped, whereas two categories were selected for potassium excretion, as the association was linear in analyses. We also completed a stratified subgroup analysis of the association between sodium excretion and outcomes, stratified by groups of urinary potassium excretion and mAHEI divided at the tertiles. We tested for interactions between sodium excretion and potassium excretion and between sodium excretion and mAHEI score. As the association of sodium excretion and composite outcome was non-linear, we tested for interaction above and below median sodium excretion (ie, four tests for interaction), modelling the log hazard ratio as a piecewise linear function of sodium excretion; two connected lines above and below the median sodium excretion with different slopes could interact with the potassium and mAHEI subgroups. We completed an analysis of “usual” sodium intake, based on repeated measurement in 494 participants, employing methodology used in previous publications.\textsuperscript{25}

To minimise the potential for reverse causation, we conducted sensitivity analyses that excluded participants with a final follow-up or outcome event within the first three years of baseline. To assess the changes in hazard for the primary outcome with simultaneous changes in sodium and potassium excretion, the hazard ratio was modelled using a two dimensional natural cubic spline. For each of sodium and potassium, two knot points were selected at the tertiles of the variable’s sample distribution. This resulted in a grid of four knot points on a two dimensional (sodium and potassium) spline surface. The Cox proportional hazard model then included (along with other covariates) natural cubic spline variables for both sodium and potassium separately and for interaction terms between these spline variables. To visualise the two dimensional change in hazard, and the putative interaction effect between sodium and potassium, estimated hazard ratios were plotted for each observed pair sodium and potassium values in the analysis dataset (a heat map). The reference for plotted hazard ratios was the hazard at the median levels of sodium (4.70 g) and potassium (2.1 g).

All analyses were conducted using R Version 3.4.4 (R Foundation for Statistical Computing, Vienna, Austria).

Patient and public involvement
No patients were involved in setting the research question or outcome measures, or in the design and implementation of the study.

Results
In total, 103 570 participants were included, of whom 41.8% were from China. In the overall population, mean estimated 24 hour sodium excretion was 4.9 g and mean estimated potassium excretion 2.1 g. (see supplementary table 1) After a median duration of 8.2 years (interquartile range 6.0-9.4) there were 4524 deaths, 4889 with a major cardiovascular event (1893 myocardial infarctions, 2526 strokes, 534 heart failure, and 1293 cardiovascular deaths), and 7884 had either a cardiovascular event or died, and 3263 had a diagnosis of cancer.

Mortality and cardiovascular events
Estimated sodium excretion
Compared with estimated sodium excretion of 4.00 to 4.99 g/day (the reference group), higher estimated sodium excretion (>7 g/day) was associated with a greater risk of the primary composite outcome (hazard ratio 1.23, 95% confidence interval 1.12 to 1.34), all cause mortality (1.36, 1.20 to 1.53),
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major cardiovascular events (1.20, 1.08 to 1.34), cardiovascular death (odds ratio 1.49, 95% confidence interval 1.21 to 1.84), and fatal stroke (hazard ratio 1.76, 95% confidence interval 1.28 to 2.41) on multivariable analysis (table 1, supplementary table 2, fig 1).

Compared with estimated sodium excretion of 4.00 to 4.99 g/day, lower estimated sodium excretion (<3 g/day) was associated with a greater risk of the primary composite outcome (hazard ratio 1.19, 95% confidence interval 1.09 to 1.30), all cause mortality (1.26, 1.12 to 1.41), major cardiovascular events (1.19, 1.06 to 1.33), cardiovascular death (1.35, 1.09 to 1.69), and stroke (1.24, 1.05 to 1.46); supplementary table 2. No association was found between sodium excretion and cancer or cancer mortality (see supplementary table 2).

**Estimated potassium excretion**

Compared with estimated potassium excretion of <1.5 g/day, higher estimated potassium excretion was associated with a lower risk of mortality and cardiovascular events (0.83, 0.73 to 0.94 for >3 g/day) on multivariable analysis (fig 2, supplementary table 3). This associated reduction in hazard was greater for mortality risk (0.71, 0.60 to 0.85) than for major cardiovascular events (0.87, 0.75 to 1.02). No apparent association was found between potassium excretion and cancer or cancer mortality (see supplementary table 3).

**Combined sodium and potassium excretion**

Table 2 shows the association of sodium excretion with the primary composite outcome, categorised by joint sodium and potassium excretion (six categories). The lowest risk of death and cardiovascular events occurred in the group with moderate sodium excretion (3.5 g/day) and higher potassium excretion (21.9% of cohort). Compared with this reference group, the combinations of low potassium with low sodium excretion (hazard ratio 1.23, 95% confidence interval 1.11 to 1.37; 7.6% of cohort) and low potassium with high sodium excretion (1.21, 1.11 to 1.32; 13.8% of cohort) were associated with the highest risk, followed by high sodium excretion (1.10, 1.02 to 1.18 (29.6% of cohort)) and low sodium excretion (1.19, 1.02 to 1.38; 3.3% of cohort) among those with potassium excretion above the median. Moderate sodium excretion with low potassium excretion was associated with an increased risk of the primary outcome (1.10, 1.01 to 1.19; 24.0% of cohort), compared with reference. Few participants (0.002%) had both very low sodium excretion (<2 g/day) and high potassium excretion (>3.5 g/day)—a target recommended by nutrition guidelines—so that the risk in this subgroup could not be estimated reliably (0.2% for potassium >2.63 g/day, assuming that about 75% of potassium is excreted in the urine, and sodium excretion <2.3 g/day). Figure 3 reports results of a heat map that combines multivariable cubic splines of the association of both sodium excretion and potassium excretion with the primary outcome, with median excretion as reference category. Table 3 and figure 4 show a stratified analysis by thirds of potassium excretion, supporting a lower relative risk associated with higher sodium excretion among those with higher potassium excretion. Test for interaction of higher sodium excretion (above the median) and potassium excretion was significant (P=0.007) but for lower sodium excretion and potassium excretion was not significant (P=0.94).

**Sodium excretion and diet quality (mAHEI)**

Figure 5 and table 3 report a stratified analysis by mAHEI thirds, which suggest a lower magnitude of risk for the association of sodium excretion with mortality and cardiovascular events among those with higher diet quality scores, although P for interactions were not significant (P=0.87 for below median excretion and P=0.30 above median excretion).

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**Table 1 | Association of estimated urinary sodium excretion with mortality and cardiovascular events. Values are hazard ratios (95% confidence intervals)**

<table>
<thead>
<tr>
<th>Variables</th>
<th>Estimated sodium excretion (g/day)</th>
</tr>
</thead>
<tbody>
<tr>
<td></td>
<td>≤3 (n=1102)</td>
</tr>
<tr>
<td>No deaths and cardiovascular events</td>
<td>949</td>
</tr>
<tr>
<td>Univariate*</td>
<td>1.23 (1.14 to 1.34)</td>
</tr>
<tr>
<td>Multivariable (primary)†</td>
<td>1.19 (1.09 to 1.30)</td>
</tr>
<tr>
<td>Multivariable (+diet)</td>
<td>1.16 (1.05 to 1.27)</td>
</tr>
<tr>
<td>Multivariable (+blood pressure/heart rate)</td>
<td>1.15 (1.04 to 1.28)</td>
</tr>
<tr>
<td>Multivariable (excluding CVD, diabetes, smoker, and cancer)§</td>
<td>1.16 (1.03 to 1.31)</td>
</tr>
<tr>
<td>Multivariable (excluding events in first 3 years)</td>
<td>1.20 (1.07 to 1.33)</td>
</tr>
<tr>
<td>Multivariable (primary+lipids)</td>
<td>1.17 (1.06 to 1.29)</td>
</tr>
</tbody>
</table>

**Section 1.**

CVD=cardiovascular disease. Major cardiovascular events include cardiovascular mortality, myocardial infarction, stroke, and heart failure.

Primary model includes the following covariates at baseline: age, sex, education, alcohol intake, diabetes mellitus, body mass index, a history of cardiovascular events, cancer and chronic obstructive pulmonary disease, cardiovascular drugs at baseline, HIV, tuberculosis, physical activity level, and smoking status and dosage. Dietary variables: addition of caloric intake, potassium intake, waist hip ratio, and modified alternative healthy eating index score. Blood pressure variables: baseline systolic blood pressure and history of hypertension. Low density lipoprotein cholesterol-high density lipoprotein cholesterol ratio available in 88% of cohort.

*Univariate analysis using Cox proportional hazards model with a random effect for study centre (to address clustering of data) includes 103 200 participants with follow-up data.

†Includes 90% of all participants in univariate model, 10% of participants had missing data for at least one covariate, we did not undertake imputation for missing data.

‡Adjusted for variables in primary model.

§Excludes those with baseline history of cardiovascular disease, diabetes mellitus, current smoking, and cancer (history of or within first year of follow-up).

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Fig 1 | Restricted cubic spline plot of association of estimated 24 hour urinary sodium excretion with composite of all cause mortality and major cardiovascular events, mortality, and major cardiovascular events (composite of cardiovascular death or myocardial infarction or stroke or heart failure). Plots adjusted for age (included as spline function), sex, education, current and former alcohol intake (units weekly), diabetes mellitus, body mass index, physical activity, history of cardiovascular events, use of cardiovascular drugs (blood pressure lowering, statins, or antidiabetics), history of tuberculosis, cancer, HIV, and current and former smoking. Dashed lines indicate 95% confidence intervals. Median intake is reference standard (4.9 g/day). Salt approximates 2.5×sodium g/day. Spline curve was truncated at 10 g/day in all plots.
Fig 2 | Restricted cubic spline plot of association of estimated 24 hour urinary potassium excretion with composite of all cause mortality and major cardiovascular events, mortality, and major cardiovascular events (composite of cardiovascular death, myocardial infarction, stroke, and heart failure). Plots adjusted for age (included as spline function), sex, education, current and former alcohol intake (units weekly), diabetes mellitus, body mass index, physical activity, history of cardiovascular events, use of cardiovascular drugs (blood pressure lowering, statins, or antidiabetics), history of tuberculosis, cancer, HIV, and current and former smoking. Spline was truncated at 4 g/day in all plots.
Sensitivity analyses
Exclusion of participants who had events in the first three years of follow-up, and excluding those with baseline cardiovascular disease, cancer (or cancer in first year of follow-up, diabetes, and current smoking), did not materially affect findings. Table 1 Supplementary figure 1 reports the association of usual 24 hour urinary sodium excretion and clinical outcomes.

Discussion
In this large international prospective cohort study, we investigated the joint association of estimated sodium and potassium urinary excretion (surrogates for intake) and clinical outcomes over a median follow-up of 8.2 years. Overall, the lowest risk of death and adverse clinical events was seen in those with estimated sodium excretion between 3 and 5 g/day, and with the highest potassium excretions (see fig 1 and table 2). Both higher and lower levels of estimated sodium excretion were associated with higher cardiovascular risk, thereby describing a J-shaped curve for association, whereas the association of potassium excretion and mortality or cardiovascular risk was inverse and linear. Higher potassium intake attenuated the increased cardiovascular risk associated with high sodium intake, and the association of high sodium excretion with cardiovascular risk was most prominent in those with low potassium intake.

Comparison with other studies
Potassium, the most abundant intracellular cation, and sodium, the most abundant extracellular cation, are inextricably linked, with their exchange required for membrane potential of cells (Na-K-ATPase) and sodium, the most abundant extracellular cation,

<table>
<thead>
<tr>
<th>Examined markers</th>
<th>Estimated sodium excretion (g/day)</th>
<th>Estimated potassium excretion (g/day)</th>
</tr>
</thead>
<tbody>
<tr>
<td>Hazard ratio</td>
<td>1.23 (1.11 to 1.37)</td>
<td>1.10 (1.01 to 1.19)</td>
</tr>
<tr>
<td>Event proportion</td>
<td>716/7582 (9.4%)</td>
<td>1924/24741 (7.8%)</td>
</tr>
<tr>
<td>Hazard ratio</td>
<td>1.19 (1.02 to 1.38)</td>
<td>1.00 (reference)</td>
</tr>
<tr>
<td>Event proportion</td>
<td>233/3420 (6.8%)</td>
<td>1454/22 688 (6.4%)</td>
</tr>
</tbody>
</table>

Adjusted for age (including as spline function), sex, education, current and former alcohol intake (units weekly), diabetes mellitus, body mass index, physical activity, history of cardiovascular events, use of cardiovascular drugs (blood pressure lowering, statins, or antidiabetics), history of tuberculosis, cancer, HIV, and current and former smoking.

*Event proportion for composite outcome of major cardiovascular events or mortality.

Fig 3 | Heat map of risk for composite of cardiovascular events or death showing lowest risk in region of moderate sodium intake 3.5 g/day and highest potassium intake and highest risk in region of extremes of sodium excretion and low potassium excretion. The reference hazard for these hazard ratios was set at a value of sodium daily excretion/intake of 5.00 g and potassium daily excretion/intake of 2.25 g (median excretion of sodium and potassium), marked as X. The overlaid lines represent joint distribution quartiles; each region contains a quarter of the analysed participants. r=0.34

Table 2 | Association of joint urinary sodium and potassium excretion with mortality and cardiovascular events

<table>
<thead>
<tr>
<th>Estimated potassium excretion (g/day)</th>
<th>Estimated sodium excretion (g/day)</th>
</tr>
</thead>
<tbody>
<tr>
<td>Median, 2.1 g/day:</td>
<td></td>
</tr>
<tr>
<td>Hazard ratio 95% CI</td>
<td>1.23 (1.11 to 1.37)</td>
</tr>
<tr>
<td>Event proportion</td>
<td>716/7582 (9.4%)</td>
</tr>
<tr>
<td>Median, 2.1 g/day:</td>
<td></td>
</tr>
<tr>
<td>Hazard ratio 95% CI</td>
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</tr>
<tr>
<td>Event proportion</td>
<td>233/3420 (6.8%)</td>
</tr>
</tbody>
</table>

Comparison with other studies
Potassium, the most abundant intracellular cation, and sodium, the most abundant extracellular cation, are inextricably linked, with their exchange required for membrane potential of cells (Na-K-ATPase) and their inter-reliance on diet and kidneys to maintain homeostasis. We found that the association of higher sodium excretion with cardiovascular risk was lower among those with higher urinary potassium excretion. While previous studies have reported that higher potassium intake diminishes the association of higher sodium intake with increased blood pressure, no study has been sufficiently large to show a statistically significant modifying effect of sodium and potassium intake with cardiovascular events and mortality. In our analyses, extremes of sodium excretion (both high and low) in a setting of low potassium excretion are associated with the highest risk of death and cardiovascular events (see fig 1 and table 2). Low potassium intake not only results in potassium retention by the kidneys, but also induces sodium retention. Renal conservation of sodium in the setting of low potassium diets is mediated through WNK (“with no lysine(K)”) kinases activating the thiazide sensitive sodium-chloride cotransporter, leading to greater increases in blood pressure with increasing sodium intake. Low potassium intake might also reflect poorer diet quality, particularly lower intake of fruits and vegetables, and the association of sodium intake with cardiovascular risk might be confounded by foods that independently affect cardiovascular risk. Our subgroup analysis by diet quality (modified alternative healthy eating index, mAHEI) would lend some support to this contention, whereby the magnitude of association of sodium excretion with cardiovascular risk appeared...
lower among those consuming the highest third of diet quality, although formal tests of interaction were not statistically significant (Table 3). Therefore, higher potassium intake could directly affect cardiovascular and mortality risks or is a marker of increased intake of healthier food items that are rich in potassium (eg, fruit, vegetables, nuts), or a combination of both. The DASH-Sodium trial reported a lower antihypertensive effect of reducing sodium intake, in the setting of higher potassium intake, supporting the contention that the association of higher sodium intake with cardiovascular health might be modified by potassium intake. A small cluster randomised controlled trial (five centres), which replaced table salt with low sodium salt substitute (partially replaced with potassium chloride) reported a reduced cardiovascular risk in those centres randomised to salt substitution, although findings were not conclusive owing to methodological limitations of the trial. An ongoing large cluster randomised controlled trial in China, evaluating the use of salt substitutes (potassium chloride substituted for sodium chloride) is expected to provide a more definitive answer. However, these trials evaluate simultaneous increased potassium intake with a reduction of high sodium intake to moderate intake levels, and it is unlikely to inform whether a very low sodium intake (<2.0 g/day) is beneficial or harmful. Finally, increased cardiovascular risk might also be mediated through activation of numerous adaptive renal and neuroendocrine mechanisms to maintain balance of both electrolytes, in response to low potassium and low dietary sodium, especially activation of the renin-aldosterone-angiotensin system, the activation of which is known to increase cardiovascular risk. Collectively, these data suggest that the health impact of these two cations cannot be easily assessed separately. However, our findings would argue against an isolated use of a sodium:potassium ratio, because of the J-shaped association of sodium intake and clinical outcomes and the ratio does not take account of absolute intake levels. For example, the combination of moderate intake of sodium and potassium is associated with lower cardiovascular risk than combined low intake of both sodium and potassium, despite the same sodium:potassium ratio.

**Public health implications**

Guidelines for dietary recommendations, targeting the entire population, should be both feasible and based on clear high quality evidence of improved health outcomes. Our results suggest that a combined strategy of low sodium intake (<2.0 g/day), while simultaneously increasing potassium intake to >3.5 g/day, is unrealistic, as only 0.002% of the population were in this range (0.2% for potassium excretion of 2.62 g/day, if we assume that about 75% of potassium is excreted in urine). These estimates are consistent with reports from studies in the United Kingdom (0.1%), United States (0.3%), Mexico (0.15%), and France (0.5%). In each of these studies, estimates of sodium intake were based on questionnaire, which is known to underestimate intake, meaning that the proportion of the population achieving actual target may be lower than reported. This finding also meant that we were unable to evaluate the association of these joint dietary intake levels with health outcomes, as numbers of participants were insufficient despite the large sample size. We observed a linear relation between increased sodium and potassium intake, a consistent finding of other studies. These findings question the feasibility of increasing potassium intake while achieving very low sodium intake of populations through modifying diets. While small phase II clinical trials, such as DASH-Sodium, have achieved these targets in the short term, they required complete control of dietary intake, which is impractical in free living populations. In contrast, the PREDMED study, which promoted increased intake of potassium containing foods (fruit, vegetables, nuts) with greater adherence to the Mediterranean diet, without focusing on low sodium intake, reported a statistically significant reduction in cardiovascular disease and mortality, and adherence to a Mediterranean diet has

### Table 3 | Association of estimated urinary sodium excretion with mortality and cardiovascular events (subgroup analysis). Values are hazard ratio (95% confidence interval)

<table>
<thead>
<tr>
<th>Estimated sodium excretion (g/day)</th>
<th>&lt;1.8 (n=11 002)</th>
<th>1.8-2.3 (n=21 417)</th>
<th>2.3-3.99 (n=26 012)</th>
<th>4.4-5.99 (n=26 012)</th>
<th>6-6.99 (n=12 458)</th>
<th>≥7 (n=11 218)</th>
<th>P for interaction</th>
</tr>
</thead>
<tbody>
<tr>
<td>Potassium excretion (g/day):</td>
<td>1.14 (1.00 to 1.30)</td>
<td>1.16 (0.97 to 1.38)</td>
<td>1.12 (0.91 to 1.39)</td>
<td>1.24 (1.06 to 1.46)</td>
<td>1.29 (1.10 to 1.52)</td>
<td>1.33 (1.10 to 1.62)</td>
<td>1.00 (1.00 to 1.00)</td>
</tr>
<tr>
<td>1.00 (0.91 to 1.14)</td>
<td>1.11 (0.97 to 1.26)</td>
<td>1.02 (0.88 to 1.18)</td>
<td>1.10 (0.97 to 1.25)</td>
<td>1.09 (0.96 to 1.25)</td>
<td>1.03 (0.90 to 1.17)</td>
<td>1.05 (0.93 to 1.19)</td>
<td></td>
</tr>
<tr>
<td>1.02 (0.89 to 1.16)</td>
<td>1.06 (1.03 to 1.19)</td>
<td>1.01 (0.88 to 1.16)</td>
<td>1.00 (0.96 to 1.11)</td>
<td>1.00 (0.88 to 1.14)</td>
<td>0.99 (0.85 to 1.14)</td>
<td>1.06 (0.91 to 1.23)</td>
<td></td>
</tr>
<tr>
<td>1.15 (0.97 to 1.36)</td>
<td>1.17 (1.01 to 1.36)</td>
<td>1.01 (0.88 to 1.16)</td>
<td>1.11 (0.96 to 1.27)</td>
<td>1.22 (1.05 to 1.43)</td>
<td>1.27 (1.04 to 1.54)</td>
<td>1.00 (1.00 to 1.00)</td>
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<tr>
<td>1.27 (1.04 to 1.54)</td>
<td>1.47 (1.26 to 1.71)</td>
<td>1.12 (1.01 to 1.25)</td>
<td>1.07 (0.92 to 1.25)</td>
<td>1.11 (0.98 to 1.25)</td>
<td>1.00 (1.00 to 1.00)</td>
<td>1.00 (1.00 to 1.00)</td>
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</tbody>
</table>

Modified alternative healthy eating index score:

| <31.3 | 1.18 (1.01 to 1.38) | 1.10 (0.97 to 1.25) | 1.00 (0.96 to 1.25) | 1.09 (0.96 to 1.25) | 1.03 (0.90 to 1.17) | 1.00 (0.88 to 1.14) | 0.99 (0.85 to 1.14) | 1.06 (0.91 to 1.23) |
| 31.3-38.4 | 1.24 (1.06 to 1.46) | 1.09 (0.96 to 1.25) | 1.00 (0.96 to 1.25) | 1.14 (1.01 to 1.30) | 1.12 (0.97 to 1.30) | 1.22 (1.05 to 1.43) | 1.27 (1.04 to 1.54) | 1.00 (1.00 to 1.00) |
| ≥38.4 | 1.09 (0.92 to 1.29) | 1.03 (0.90 to 1.17) | 1.00 (0.96 to 1.11) | 1.00 (0.88 to 1.14) | 0.99 (0.85 to 1.14) | 1.06 (0.91 to 1.23) | 1.00 (1.00 to 1.00) | 1.00 (1.00 to 1.00) |

Adjusted for age (included as spline function), sex, education, current and former alcohol intake (units per week), diabetes mellitus, body mass index, physical activity, history of cardiovascular events, use of cardiovascular medications (blood pressure lowering, statins, or antidiabetics), history of tuberculosis, cancer, HIV, and current and former smoking.
Fig 4 | Association of urinary sodium excretion within tertiles of urinary potassium excretion. P for interaction is significant (P=0.007) for urinary potassium excretion×urinary sodium excretion above median intake, with lower magnitude of association with higher urinary potassium excretion. Plots adjusted for age (included as spline function), sex, education, current and former alcohol intake (units weekly), diabetes mellitus, body mass index, physical activity, history of cardiovascular events, use of cardiovascular drugs (blood pressure lowering, statins, or antidiabetics), history of tuberculosis, cancer, HIV, and current and former smoking.
Fig 5 | Association of urinary sodium excretion within tertiles of modified alternative healthy eating index (mAHEI) score. P for interactions not significant. Plots adjusted for age (included as spline function), sex, education, current and former alcohol intake (units weekly), diabetes mellitus, body mass index, physical activity, history of cardiovascular events, use of cardiovascular drugs (blood pressure lowering, statins, or antidiabetics), history of tuberculosis, cancer, HIV, and current and former smoking.
In another study,48 which used an appropriate fasting urine sample, best reflecting basal excretion. Kawasaki formula was developed and validated for a expected to produce biased over-estimates since the fasting spot urine sample was used, which would be estimates.31 We chose this practical approach to the correlation we report for formula derived glucose to glycated haemoglobin (HbA1c), is similar to 24 hour ambulatory blood pressure and of fasting glucose, and cholesterol to clinical outcomes. For example, the correlation of office blood pressure and potassium intake to long term changes in dietary intake over time, there is stability in mean intake of countries,6 51 especially those with mean sodium intake in the moderate range (3-5 g/day), whereas some countries with high mean intake (>5 g/day) have reported reductions in mean intake levels (eg, China, Finland),52 which means that interval changes in sodium intake are less likely to influence our findings in the low to moderate range of sodium intake but could reduce the magnitude of association in the higher sodium intake range.

Strengths and limitations of this study
Our estimate of sodium intake was based on a baseline measurement of sodium intake, derived from fasting morning urine samples, rather than repeated 24 hour urinary collections, which would be considered the reference standard for estimating usual sodium and potassium intake. However, our approach has been validated in an international study, against actual 24 hour urine estimates of sodium and potassium excretion.31 Other formula based approaches have been developed for estimating 24 hour urinary sodium and potassium excretion, and one study reported that the Kawasaki formula was associated with the largest bias compared with actual 24 hour urine collections. In that study by Cogswell et al,47 however, a non-fasting spot urine sample was used, which would be expected to produce biased over-estimates since the Kawasaki formula was developed and validated for a fasting urine sample, best reflecting basal excretion. In another study,48 which used an appropriate fasting sample, the Kawasaki formula approach was reported to be associated with least biased estimates (versus other formula) compared with actual 24 hour urine. Further evidence of the construct validity of our approach is the relation of estimated 24 hour urine estimates of sodium excretion to blood pressure (2.11/0.78 mm Hg/g increase in estimated 24 hour urinary sodium excretion), which are consistent with estimates reported in randomised controlled trials of sodium reduction.25

Relating a single time-point estimate of sodium and potassium intake to long term follow-up of death and cardiovascular events might be problematic if there are interval changes in dietary intake over time, and where repeated dietary measures are required to reduce the risk of regression dilution bias. Our approach is consistent with many other large epidemiological studies of sodium and potassium intake and analogous to epidemiological studies relating blood pressure, glucose, and cholesterol to clinical outcomes. For example, the correlation of office blood pressure and 24 hour ambulatory blood pressure and of fasting glucose to glycated haemoglobin (HbA1c), is similar to the correlation we report for formula derived estimates.31 We chose this practical approach to estimating population level sodium and potassium intake in the PURE study, which includes a large representative population of participants from high, middle, and low income countries. We deemed actual 24 hour collections to be unfeasible in many low income settings and expected to have a greater risk selection bias due to non-completion of 24 hour collections. For example, in a recent NHANES study,49 completion rates for 24 hour urine samples were about 75%, and completers differed from non-completed for age, sex, ethnicity, body mass index, and hypertension status. For these reasons, WHO suggests50 the use of formula derived estimates of sodium and potassium intake in population level studies monitoring intakes over time. Among studies reporting on temporal changes of sodium intake over time, there is stability in mean intake of countries.6 51 especially those with mean sodium intake in the moderate range (3-5 g/day), whereas some countries with high mean intake (>5 g/day) have reported reductions in mean intake levels (eg, China, Finland),52 which means that interval changes in sodium intake are less likely to influence our findings in the low to moderate range of sodium intake but could reduce the magnitude of association in the higher sodium intake range.

Reverse causation has been a criticism of prospective cohort studies evaluating the association of diet and clinical outcomes, which may also occur when people reduce their sodium intake because of illness or based on medical recommendations. Alternatively, it may occur when people with a history of cardiovascular disease or risk factors, who are already at increased cardiovascular risk, reduce their sodium intake on medical recommendations. In the PURE study, a few participants had previous cardiovascular disease, and our findings were not materially altered with their exclusion. Moreover, in the current analyses the exclusion of participants with cancer or diabetes or who were current smokers or those who had events in the first three years of follow-up did not materially alter findings. Additionally, we did not find an increased risk of death due to cancer with low sodium intake. Finally, given the observational nature of the study, residual confounding is a potential limitation. However, we adjusted for all major confounders of the association between sodium and potassium intake and cardiovascular events. Strengths of our study include the large representative international population in the PURE study, completeness of follow-up, rigorous approach to measurement of baseline variables, and adjudicated outcome clinical events.

Conclusions
Our findings suggest that a simultaneous target of low sodium (<2 g/day) with high potassium intake (>3.5 g/day) in the population is extremely uncommon. The combination of moderate sodium intake (3-5 g/day) with high potassium intake is associated with the lowest risk of mortality and cardiovascular events, whereas extremes of sodium intake combined with low potassium urinary excretion were associated with the highest cardiovascular risk. Our data support population wide increases in dietary potassium intake, with a population specific (ie, sodium intakes >5 g/
day) reduction of sodium intake, embedded within an overall healthy dietary pattern.

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Funding: The funders of the study had no role in its design or conduct, in the collection, analysis, or interpretation of the data, or in the writing of the manuscript. The current analyses were supported by funding from the European Research Council (COSIP grant, 640580) and Heart and Stroke Foundation of Ontario. SY is supported by the Mary W Burke endowed chair of the Heart and Stroke Foundation of Ontario. The PURE study is an investigator initiated study that is funded by the Population Health Research Institute, the Canadian Institutes of Health Research, Heart and Stroke Foundation of Ontario, Support from CIHR’s Strategy for Patient-Oriented Research, through the Ontario SPOR Support Unit, as well as the Ontario Ministry of Health and Long-Term Care and through unrestricted grants from several pharmaceutical companies (with major contributions from AstraZeneca (Canada), Sanofi-Aventis (France and Canada), Boehringer Ingelheim (Germany and Canada), Servier, and GSK), and additional contributions from Novartis and King Pharma and from various national or local organisations in participating countries. These include Argentina: Fundacion ECLA; Bangladesh: Independent University, Bangladesh and Mitra and Associates; Brazil: Unilever Health Institute, Brazil; Canada: Public Health Agency of Canada and Champlain Cardiovascular Disease Prevention Network; Canadian Institutes of Health Research (Catalyst Grant: eHealth Innovations grant No: 126524); Chile: Universidad de La Frontera (Internal Registry Di13-PE11); China: National Center for Cardiovascular Diseases; Colombia: Colciencias, grant No 6566-04-18062; India: Indian Council of Medical Research; Malaysia: Ministry of Science, Technology and Innovation of Malaysia grant No 100-IRDC/BIOETEK/16/6/21 (13/2007), grant No 07-05-IFN-BPH O10, Ministry of Higher Education of Malaysia grant No 600-RMI/ LRS5/3 (1/2011), Universiti Teknologi MARA, Faculty of Medicine, Universiti Kebangsaan Malaysia, Kuala Lumpur Malaysia, CRIM, University Kebangsaan Malaysia; Occupied Palestinian Territory: The United Nations Relief and Works Agency for Palestine Refugees in the Near East (UNRWA), occupied Palestinian territory, International Development Research Centre (IDRC), Canada, Penu: National Heart, Lung and Blood Institute (HHSN26820090003C), National Cancer Institute (1F22CA131723), Wellcome Trust (039494/2/14/2), Philippines: Philippine Council for Health Research & Development (PHCHRDRD), Poland: Polish Ministry of Science and Higher Education grant No 290/W-PURE/2008/0, Wroclaw Medical University, Wroclaw Medical University, statutory activity Saudi Arabia: Saudi Heart Association, The Deanship of Scientific Research at King Saud University, Riyadh, Saudi Arabia (Research group No: RG -1436-013), South Africa: The North-West University, SANPAD (SA and Netherlands Programme for Alternative Development), National Research Foundation, Medical Research Council of SA, The SA Sugar Association (SASA), Faculty of Community and Health Sciences (UWC), Sweden: grants from the Swedish State under the agreement concerning research and education of doctors; the Swedish Heart and Lung Foundation; the Swedish Research Council; the Swedish Council for Health, Working Life and Welfare, King Gustaf V:s and Queen Victoria Freemason’s Foundation, AFA Insurance, Tanzania: Pamoja Tunaweza Health Research Centre (Tanzania), Queen’s University, Department of Medicine, Turkey: Metabolic Syndrome Society, Astra Zeneca, Turkey; UAE: Sheikh Hamdan Bin Rashid Al Maktoum Award for Medical Sciences, and Dubai Health Authority, Dubai UAE.

Competing interests: All authors have completed the ICMJE uniform disclosure form at www.icmje.org/coiDisclosure.pdf and declare: no support from any additional organisation for the submitted work. A detailed list of funders is provided in the supplementary appendix. The authors have no financial relationships with any organisations, or other relationships or activities, that might have influenced the submitted work in the previous three years.

Ethical approval: This study was approved by the research ethics committees at all participating centres and at Hamilton Health Sciences, Hamilton, Ontario, Canada.

Data sharing: No additional data available.

Transparency: The lead author (MOD) affirms that the manuscript is an honest, accurate, and transparent account of the study being reported; that no important aspects of the study have been omitted and that any discrepancies from the study as planned have been explained.

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the BMJ | BMJ: first published as 10.1136/bmj.l772 on 13 March 2019. Downloaded from http://www.bmj.com on 2 April 2019 by guest. Protected by copyright.


**Supplementary information:** additional material